

\mathcal{I} -Cesàro Summability of a Sequence of Order α of Random Variables in Probability

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Abstract

In this paper, we define four types of convergence of a sequence of random variables, namely, \mathcal{I} -statistical convergence of order α , \mathcal{I} -lacunary statistical convergence of order α , strongly \mathcal{I} -lacunary convergence of order α and strongly \mathcal{I} -Cesàro summability of order α in probability where $0 < \alpha < 1$. We establish the connection between these notions.

1. Introduction and background

Theory of statistical convergence was firstly originated by Fast [1]. After Fridy [2] and Šalát [3] statistical convergence became a notable topic in summability theory. Lacunary statistical convergence was defined by using lacunary sequences in [4]. \mathcal{I} -convergence was firstly considered by Kostyrko et al. [5]. Also, Das et al. [6] gave new definitions by using ideal, such as \mathcal{I} -statistical convergence, \mathcal{I} -lacunary statistical convergence. Ulusu et al. [7] also studied asymptotically \mathcal{I} -Cesàro equivalence of sequences of sets.

Statistical convergence of order α ($0 < \alpha < 1$) was introduced using the notion of natural density of order α where n is replaced by n^α in [8]. This new type convergence was different in many ways from statistical convergence. Lacunary statistical convergence of order α is studied by Sengöl and M. Et [9], \mathcal{I} -statistical and \mathcal{I} -lacunary statistical convergence of order α is studied by Das and Savas [10].

In probability theory, if for $n > 0$, a random variable X_n given on space S , a probability function $P : X \rightarrow \mathbb{R}$, then we say that $X_1, X_2, \dots, X_n, \dots$ is a sequence of random variables and it is demonstrated by $\{X_n\}_{n \in \mathbb{N}}$.

It is important that if there exists $c \in \mathbb{R}$ for which $P(|X - c| < \varepsilon) = 1$, where $\varepsilon > 0$ is sufficiently small, that is, it means that values of X lie in a very small neighbourhood of c .

New concepts have begun to be studied in probability theory by Das et al. [6], and others ([11]-[15]).

2. Main results

Definition 2.1. $\{X_k\}_{k \in \mathbb{N}}$ is said to be \mathcal{I} -statistically convergent of order α in probability to a random variable X if for any $\varepsilon, \delta, \gamma > 0$

$$\left\{ n \in \mathbb{N} : \frac{1}{n^\alpha} |\{k \leq n : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \geq \gamma \right\} \in \mathcal{I},$$

and demonstrated by $X_k \xrightarrow{PS(\mathcal{I})^\alpha} X$.

Definition 2.2. $\{X_n\}_{n \in \mathbb{N}}$ is said to be \mathcal{I} -lacunary statistically convergent of order α in probability to a random variable X if for any $\varepsilon, \delta, \gamma > 0$

$$\left\{ r \in \mathbb{N} : \frac{1}{h_r^\alpha} |\{k \in I_r : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \geq \gamma \right\} \in \mathcal{I},$$

and it is demonstrated by $X_k \xrightarrow{PS_\theta(\mathcal{I})^\alpha} X$.

Definition 2.3. $\{X_k\}_{k \in \mathbb{N}}$ is said to be strongly \mathcal{I} -lacunary convergent or $PV_{\theta}(\mathcal{I})$ -convergent of order α in probability to a random variable X if for every $\varepsilon, \delta > 0$,

$$\left\{ r \in \mathbb{N} : \frac{1}{h_r^\alpha} \sum_{k \in I_r} P(|X_k - X| \geq \varepsilon) \geq \delta \right\} \in \mathcal{I},$$

and it is demonstrated by $X_k \xrightarrow{PV_{\theta}(\mathcal{I})^\alpha} X$.

Definition 2.4. $\{X_k\}_{k \in \mathbb{N}}$ is said to be strongly \mathcal{I} -Cesàro summable of order α in probability to a random variable X if for every $\varepsilon, \delta > 0$,

$$\left\{ n \in \mathbb{N} : \frac{1}{n^\alpha} \sum_{k=1}^n P(|X_k - X| \geq \varepsilon) \geq \delta \right\} \in \mathcal{I},$$

and it is demonstrated by $X_k \xrightarrow{PC_1[\mathcal{I}]^\alpha} X$.

Theorem 2.5. If $0 < \alpha \leq \beta \leq 1$ then $PS(\mathcal{I})^\alpha \subseteq PS(\mathcal{I})^\beta$.

Proof. From the assumption, we say that

$$\frac{1}{n^\beta} |\{k \leq n : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \leq \frac{1}{n^\alpha} |\{k \leq n : P(|X_k - X| \geq \varepsilon) \geq \delta\}|$$

Hence,

$$\begin{aligned} & \left\{ n \in \mathbb{N} : \frac{1}{n^\beta} |\{k \leq n : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \geq \gamma \right\} \\ & \left\{ n \in \mathbb{N} : \frac{1}{n^\alpha} |\{k \leq n : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \geq \gamma \right\} \end{aligned}$$

for $\gamma > 0$. Therefore, we obtain $PS(\mathcal{I})^\alpha \subseteq PS(\mathcal{I})^\beta$. □

Theorem 2.6. If $\liminf_r q_r > 1$, then

$$X_k \xrightarrow{PC_1[\mathcal{I}]^\alpha} X \Rightarrow X_k \xrightarrow{PV_{\theta}(\mathcal{I})^\alpha} X.$$

Proof. If $\liminf_r q_r > 1$, there exists $\gamma > 0$ such that $q_r \geq 1 + \gamma$ for all $r \geq 1$. Since $h_r = k_r - k_{r-1}$, we have $\frac{k_r^\alpha}{h_r^\alpha} \leq \left(\frac{1+\gamma}{\gamma}\right)^\alpha$ and $\frac{k_{r-1}^\alpha}{h_r^\alpha} \leq \left(\frac{1}{\gamma}\right)^\alpha$. Let $\varepsilon > 0$ and we define set by

$$S = \left\{ k_r \in \mathbb{N} : \frac{1}{k_r^\alpha} \sum_{k=1}^{k_r} P(|X_k - X| \geq \varepsilon) < \delta \right\}.$$

Therefore, $S \in \mathcal{F}(\mathcal{I})$.

$$\begin{aligned} \frac{1}{h_r^\alpha} \sum_{k \in I_r} P(|X_k - X| \geq \varepsilon) &= \frac{1}{h_r^\alpha} \sum_{k=1}^{k_r} P(|X_k - X| \geq \varepsilon) - \frac{1}{h_r^\alpha} \sum_{k=1}^{k_{r-1}} P(|X_k - X| \geq \varepsilon) \\ &= \frac{k_r^\alpha}{h_r^\alpha} \cdot \frac{1}{k_r^\alpha} \sum_{k=1}^{k_r} P(|X_k - X| \geq \varepsilon) - \frac{k_{r-1}^\alpha}{h_r^\alpha} \cdot \frac{1}{k_{r-1}^\alpha} \sum_{k=1}^{k_{r-1}} P(|X_k - X| \geq \varepsilon) \\ &\leq \left(\frac{1+\gamma}{\gamma}\right)^\alpha \delta - \left(\frac{1}{\delta\gamma}\right)^\alpha \delta' \end{aligned}$$

for each $k_r \in S$. Choose $\eta = \left(\frac{1+\gamma}{\gamma}\right)^\alpha \delta - \left(\frac{1}{\delta\gamma}\right)^\alpha \delta'$. Therefore,

$$\left\{ r \in \mathbb{N} : \frac{1}{h_r^\alpha} \sum_{k \in I_r} P(|X_k - X| \geq \varepsilon) < \eta \right\} \in \mathcal{F}(\mathcal{I}).$$

Hence, we get $X_k \xrightarrow{PV_{\theta}(\mathcal{I})^\alpha} X$. □

Theorem 2.7. If $\{X_k\}$ is strongly \mathcal{I} -Cesàro summable of order α then, it is \mathcal{I} -statistical convergent of order α in probability to a random variable X .

Proof. Let $X_k \xrightarrow{PC_1[\mathcal{I}]^\alpha} X$, and $\epsilon > 0$ given. Then

$$\begin{aligned} \frac{1}{n^\alpha} \sum_{k=1}^n P(|X_k - X| \geq \epsilon) &\geq \frac{1}{n^\alpha} \sum_{\substack{k=1 \\ P(|X_k - X| \geq \epsilon)}}^n P(|X_k - X| \geq \epsilon) \\ &\geq \frac{\delta}{n^\alpha} \cdot |\{k \leq n : P(|X_k - X| \geq \epsilon) \geq \delta\}| \end{aligned}$$

and so

$$\frac{1}{\delta \cdot n^\alpha} \sum_{k=1}^n P(|X_k - X| \geq \epsilon) \geq \frac{1}{n^\alpha} |\{k \leq n : P(|X_k - X| \geq \epsilon) \geq \delta\}|.$$

So for a given $\tau > 0$,

$$\begin{aligned} &\{n \in \mathbb{N} : \frac{1}{n^\alpha} |\{k \leq n : P(|X_k - X| \geq \epsilon) \geq \delta\}| \geq \tau\} \\ &\subseteq \left\{ n \in \mathbb{N} : \frac{1}{n^\alpha} \sum_{k=1}^n P(|X_k - X| \geq \epsilon) \geq \delta \cdot \tau \right\} \in \mathcal{I}. \end{aligned}$$

Therefore, $X_k \xrightarrow{PS(\mathcal{I})^\alpha} X$. □

Theorem 2.8. Let a bounded $\{X_k\}$ is \mathcal{I} -statistical convergent of order α to X . Hence, it is strongly \mathcal{I} -Cesàro summable of order α to X .

Proof. Assume that $\{X_k\}$ is bounded and $X_k \xrightarrow{PS(\mathcal{I})^\alpha} X$. Since $\{X_k\}$ is bounded, we get $P(|X_k - X| > \epsilon) \leq M$ for all k . For $\epsilon > 0$, we have

$$\begin{aligned} \frac{1}{n^\alpha} \sum_{k=1}^n P(|X_k - X| \geq \epsilon) &= \frac{1}{n^\alpha} \sum_{\substack{k=1 \\ P(|X_k - X| \geq \epsilon) \geq \delta}}^n P(|X_k - X| \geq \epsilon) \\ &\quad + \frac{1}{n^\alpha} \sum_{\substack{k=1 \\ P(|X_k - X| \geq \epsilon) < \delta}}^n P(|X_k - X| \geq \epsilon) \\ &\leq \frac{1}{n^\alpha} M |\{k \leq n : P(|X_k - X| \geq \epsilon) \geq \delta\}| \\ &\quad + \frac{1}{n^\alpha} n^\alpha \delta \end{aligned}$$

Then for any $\gamma > 0$,

$$\begin{aligned} &\left\{ n \in \mathbb{N} : \frac{1}{n^\alpha} \sum_{k=1}^n P(|X_k - X| \geq \epsilon) \geq \gamma \right\} \\ &\subseteq \left\{ n \in \mathbb{N} : \frac{1}{n^\alpha} |\{k \leq n : P(|X_k - X| \geq \epsilon) \geq \delta\}| \geq \frac{\gamma}{M} \right\} \in \mathcal{I}. \end{aligned}$$

Therefore $X_k \xrightarrow{PC_1[\mathcal{I}]^\alpha} X$. □

Theorem 2.9. For $\theta = \{k_r\}$,

- (i) If $\{X_k\} \xrightarrow{PV_\theta(\mathcal{I})^\alpha} X$ then $\{X_k\} \xrightarrow{PS_\theta(\mathcal{I})^\alpha} X$, and
- (ii) $PV_\theta(\mathcal{I})^\alpha$ is proper subset of $PS_\theta(\mathcal{I})^\alpha$.

Proof. (i) Let $\epsilon, \delta > 0$ and $\{X_k\} \xrightarrow{PV_\theta(\mathcal{I})^\alpha} X$. Then, we can write

$$\begin{aligned} \frac{1}{h_r^\alpha} \sum_{k \in I_r} P(|X_k - X| \geq \epsilon) &\geq \frac{1}{h_r^\alpha} \sum_{\substack{k \in I_r \\ P(|X_k - X| \geq \epsilon) \geq \delta}} P(|X_k - X| \geq \epsilon) \\ &\geq \frac{\delta}{h_r^\alpha} \cdot |\{k \in I_r : P(|X_k - X| \geq \epsilon) \geq \delta\}|. \end{aligned}$$

Therefore

$$\frac{1}{\delta h_r^\alpha} \sum_{k \in I_r} P(|X_k - X| \geq \epsilon) \geq \frac{1}{h_r^\alpha} \cdot |\{k \in I_r : P(|X_k - X| \geq \epsilon) \geq \delta\}|.$$

which implies that for any $\gamma > 0$,

$$\begin{aligned} &\left\{ r \in \mathbb{N} : \frac{1}{h_r^\alpha} |\{k \in I_r : P(|X_k - X| \geq \epsilon) \geq \delta\}| \geq \gamma \right\} \\ &\subseteq \left\{ r \in \mathbb{N} : \frac{1}{h_r^\alpha} \sum_{k \in I_r} P(|X_k - X| \geq \epsilon) \geq \delta \gamma \right\} \in \mathcal{I}. \end{aligned}$$

Hence we get $X_k \xrightarrow{PS_{\theta}(\mathcal{I})^\alpha} X$.

(ii) Let $\{X_k\}$ be defined by

$$X_k = \begin{cases} \{-1, 1\} & , \text{ with probability } \frac{1}{2}, \text{ if } n \text{ is the first } \lceil \sqrt{h_r^\alpha} \rceil \text{ integers in the interval } I_r, \\ \{0, 1\} & , \text{ with probability } P(X_n = 0) = \left(1 - \frac{1}{n}\right) \text{ and } P(X_n = 1) = \frac{1}{n}, \\ & \text{if } n \text{ is other than the first} \\ & \lceil \sqrt{h_r^\alpha} \rceil \text{ integers in the interval } I_r. \end{cases}$$

Let $0 < \varepsilon < 1$ and $\delta < 1$. Then, we obtain

$$P(|X_k - 0| \geq \varepsilon) = \begin{cases} 1 & , \text{ if } n \text{ is the first } \lceil \sqrt{h_r^\alpha} \rceil \text{ integers in the interval } I_r, \\ \frac{1}{n} & \text{if } n \text{ is other than the first } \lceil \sqrt{h_r^\alpha} \rceil \text{ integers in the interval } I_r. \end{cases}$$

Now

$$\frac{1}{h_r^\alpha} |\{k \in I_r : P(|X_k - 0| \geq \varepsilon) \geq \delta\}| \leq \frac{\lceil \sqrt{h_r^\alpha} \rceil}{h_r^\alpha}$$

and for any $\gamma > 0$ we get

$$\left\{ r \in \mathbb{N} : \frac{1}{h_r^\alpha} |\{k \in I_r : P(|X_k - 0| \geq \varepsilon) \geq \delta\}| \geq \gamma \right\} \subseteq \left\{ r \in \mathbb{N} : \frac{\lceil \sqrt{h_r^\alpha} \rceil}{h_r^\alpha} \geq \gamma \right\}.$$

Since the set

$$\left\{ r \in \mathbb{N} : \frac{\lceil \sqrt{h_r^\alpha} \rceil}{h_r^\alpha} \geq \gamma \right\}$$

is finite and so belongs to \mathcal{I} , therefore, we obtain

$$\left\{ r \in \mathbb{N} : \frac{1}{h_r^\alpha} |\{k \in I_r : P(|X_k - 0| \geq \varepsilon) \geq \delta\}| \geq \gamma \right\} \in \mathcal{I}$$

which means that $X_k \xrightarrow{PS_{\theta}(\mathcal{I})^\alpha} 0$. Also,

$$\frac{1}{h_r^\alpha} \sum_{k \in I_r} P(|X_k - 0| \geq \varepsilon) = \frac{1}{h_r^\alpha} \cdot \frac{\lceil \sqrt{h_r^\alpha} \rceil (\lceil \sqrt{h_r^\alpha} \rceil + 1)}{2},$$

then

$$\begin{aligned} \left\{ r \in \mathbb{N} : \frac{1}{h_r} \sum_{k \in I_r} P(|X_k - 0| \geq \varepsilon) \geq \frac{1}{4} \right\} &= \left\{ r \in \mathbb{N} : \frac{\lceil \sqrt{h_r^\alpha} \rceil (\lceil \sqrt{h_r^\alpha} \rceil + 1)}{h_r} \geq \frac{1}{2} \right\} \\ &= \{m, m + 1, m + 2, \dots\} \in \mathcal{F}(\mathcal{I}) \end{aligned}$$

for some $m \in \mathbb{N}$. Hence, $X_k \xrightarrow{PS_{\theta}(\mathcal{I})^\alpha} 0$. □

Theorem 2.10. \mathcal{I} -statistical convergence in probability of order α implies \mathcal{I} -lacunary statistical convergence in probability of order α $\liminf_r q_r > 1$.

Proof. By assumption $\liminf_r q_r > 1$, then there exists a $\sigma > 0$ such that $q_r \geq 1 + \sigma$ for sufficiently large r , that is,

$$\frac{h_r}{k_r} \geq \frac{\sigma}{1 + \sigma} \Rightarrow \frac{1}{h_r^\alpha} \leq \frac{1}{k_r^\alpha} \left(\frac{1 + \sigma}{\sigma} \right)^\alpha$$

If $\{X_k\} \xrightarrow{PS(\mathcal{I})^\alpha} X$, then for $\varepsilon > 0$ and for $r > 0$, we have

$$\frac{1}{h_r^\alpha} |\{k \in I_r : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \leq \frac{1}{k_r^\alpha} \left(\frac{1 + \sigma}{\sigma} \right)^\alpha |\{k \leq k_r : P(|X_k - X| \geq \varepsilon) \geq \delta\}|$$

Then for any $\gamma > 0$, we get

$$\begin{aligned} \left\{ r \in \mathbb{N} : \frac{1}{h_r^\alpha} |\{k \in I_r : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \geq \gamma \right\} \\ \subseteq \left\{ r \in \mathbb{N} : \frac{1}{k_r^\alpha} |\{k \leq k_r : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \geq \frac{\gamma \sigma^\alpha}{(1 + \sigma)^\alpha} \right\} \in \mathcal{I}. \end{aligned}$$

□

Theorem 2.11. \mathcal{I} -lacunary statistical convergence in probability of order α implies \mathcal{I} -statistical convergence in probability of order α , $0 < \alpha < 1$, if $\sup_r \sum_{i=0}^{r-1} \frac{h_{i+1}^\alpha}{(k_{r-1})^\alpha} = B < \infty$.

Proof. Suppose that $\{X_k\} \xrightarrow{PS_\theta(\mathcal{I})^\alpha} X$, and for $\varepsilon, \delta, \gamma_1, \gamma_2 > 0$ define the sets

$$C = \left\{ r \in \mathbb{N} : \frac{1}{h_r^\alpha} |\{k \in I_r : P(|X_k - X| \geq \varepsilon) \geq \delta\}| < \gamma_1 \right\}$$

and

$$T = \left\{ n \in \mathbb{N} : \frac{1}{n^\alpha} |\{k \leq n : P(|X_k - X| \geq \varepsilon) \geq \delta\}| < \gamma_2 \right\}.$$

From our assumption we get $C \in \mathcal{F}(\mathcal{I})$. Further observe that

$$K_j = \frac{1}{h_j^\alpha} |\{k \in I_j : P(|X_k - X| \geq \varepsilon) \geq \delta\}| < \gamma_1$$

for all $j \in C$. Let $n \in \mathbb{N}$ be such that $k_{r-1} < n \leq k_r$ for some $r \in C$. Hence, we obtain

$$\begin{aligned} & \frac{1}{n^\alpha} |\{k \leq n : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \\ & \leq \frac{1}{k_{r-1}^\alpha} |\{k \leq k_r : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \\ & = \frac{1}{k_{r-1}^\alpha} |\{k \in I_1 : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \\ & \quad + \frac{1}{k_{r-1}^\alpha} |\{k \in I_2 : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \\ & \quad + \dots + \frac{1}{k_{r-1}^\alpha} |\{k \in I_r : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \\ & = \frac{k_1^\alpha}{k_{r-1}^\alpha} \frac{1}{h_1^\alpha} |\{k \in I_1 : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \\ & \quad + \frac{(k_2 - k_1)^\alpha}{k_{r-1}^\alpha} \frac{1}{h_2^\alpha} |\{k \in I_2 : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \\ & \quad + \dots + \frac{(k_r - k_{r-1})^\alpha}{k_{r-1}^\alpha} \frac{1}{h_r^\alpha} |\{k \in I_r : P(|X_k - X| \geq \varepsilon) \geq \delta\}| \\ & = \frac{k_1^\alpha}{k_{r-1}^\alpha} K_1 + \frac{(k_2 - k_1)^\alpha}{k_{r-1}^\alpha} K_2 + \dots + \frac{(k_r - k_{r-1})^\alpha}{k_{r-1}^\alpha} K_r \\ & \leq \left\{ \sup_{j \in C} K_j \right\} \sup_r \sum_{i=0}^{r-1} \frac{h_{i+1}^\alpha}{(k_{r-1})^\alpha} \\ & < \gamma_1 B. \end{aligned}$$

Choosing $\gamma_2 = \frac{\gamma_1}{B}$ and by $\bigcup \{n : k_{r-1} < n \leq k_r, r \in C\} \subset T$ where $C \in \mathcal{F}(\mathcal{I})$ Then the set T belongs to $\mathcal{F}(\mathcal{I})$ and this completes the proof. \square

References

- [1] H. Fast, *Sur la convergence statistique*, Coll. Math., **2** (1951), 241-244.
- [2] J. A. Fridy, *On statistical convergence*, Analysis, **5** (1985), 301-313.
- [3] T. Šalát, *On statistically convergent sequences of real numbers*, Math. Slovaca, **30**(2) (1980), 139-150.
- [4] J. A. Fridy, C. Orhan, *Lacunary statistical convergence*, Pacific J. Math., **160**(1) (1993), 43-51.
- [5] P. Kostyrko, T. Šalát, W. Wilczyński, *\mathcal{I} -Convergence*, Real Anal. Exchange, **26**(2) (2000), 669-686.
- [6] P. Das, E. Savaş, S. Ghosal, *On generalization of certain summability methods using ideals*, Appl. Math. Lett., **24** (2011), 1509-1514.
- [7] U. Ulusu, E. Dundar, *Asymptotically I-Cesaro equivalence of sequences of sets*, Univers. J. Math. Appl., **1**(2) (2018), 101-105.
- [8] R. Çolak, *Statistical convergence of order α* , Modern methods in analysis and its applications, Anamaya Pub., New Delhi, India, (2010), 121-129.
- [9] H. Şengöl, M. Et, *On lacunary statistical convergence of order α* , Acta Math. Sci., **34B**(2) (2014), 473-482.
- [10] P. Das, E. Savaş, *On \mathcal{I} -statistical and \mathcal{I} -lacunary statistical convergence of order α* , Bull. Iranian Math. Soc., **40**(2) (2014), 459-472.
- [11] S. Ghosal, *Statistical convergence of a sequence of random variables and limit theorems*, Appl. Math., **4**(58) (2013), 423-437.
- [12] S. Ghosal, *\mathcal{I} -statistical convergence of a sequence of random variables in probability*, Afrika Mat., <http://dx.doi.org/10.1007/s13370-013-0142-x>.
- [13] S. Ghosal, *S_λ -statistical convergence of a sequence of random variables*, J. Egypt. Math. Soc., (2014), <http://dx.doi.org/10.1016/j.joems.2014.03.007>.
- [14] S. Ghosal, *Statistical convergence of order α in probability*, Arab J. Math. Sci., **21** (2015), 253-265.
- [15] I. J. Schoenberg, *The integrability of certain functions and related summability methods*, Amer. Math. Monthly, **66** (1959), 361-375.